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# Lecture Recording

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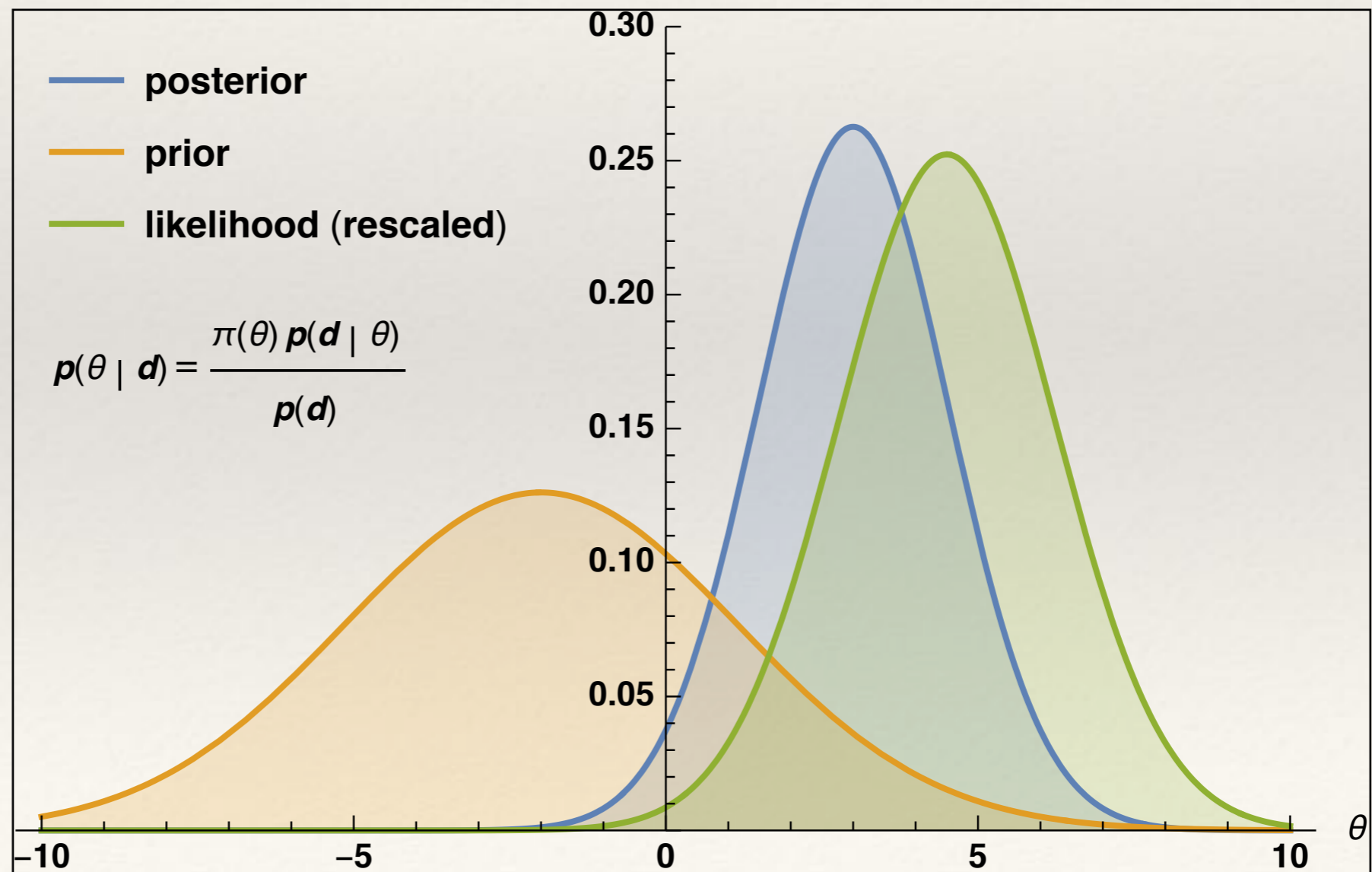
- ❖ **Note: These lectures will be recorded and posted onto the IMPRS website**
- ❖ Dear participants,
- ❖ We will record all lectures on “*Making sense of data: introduction to statistics for gravitational wave astronomy*”, including possible Q&A after the presentation, and we will make the recordings publicly available on the IMPRS lecture website at:
  - <https://imprs-gw-lectures.aei.mpg.de/2023-making-sense-of-data/>
- ❖ By participating in this Zoom meeting, you are giving your explicit consent to the recording of the lecture and the publication of the recording on the course website.

# Making sense of data: introduction to statistics for gravitational wave astronomy

## Lecture 1: introduction to random variables

*AEI IMPRS Lecture Course*

*Jonathan Gair jgair@aei.mpg.de*



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# Outline of course

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- ❖ Lectures will take place at 11am Monday, Wednesday, Thursday and Friday in the weeks beginning Nov 13<sup>th</sup>, 20<sup>th</sup> and 27<sup>th</sup> In the week beginning Dec 4<sup>th</sup> there will be lectures on Monday and Friday only. The last two lectures will take place on December 13<sup>th</sup> and 15<sup>th</sup>.
- ❖ Lectures will all take place in seminar room 0.01 at the AEI and will also be broadcast via Zoom
  - Meeting ID: 610 8405 1709
  - Meeting password: 797295
- ❖ Lecture recordings will be made available on the course website
  - <https://imprs-gw-lectures.aei.mpg.de/2023-making-sense-of-data/>

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# Outline of course

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- ❖ **Part 1 (lectures 1 to 6): Frequentist statistics and stochastic processes**
  - Random variables: definition, properties, some useful probability distributions, central limit theorem.
  - Statistics: definition, estimators, likelihood, desirable properties of estimators, Cramer-Rao bound.
  - Hypothesis testing: definition, Neyman-Pearson lemma, power and size of tests, type I and type II errors, ROC curves, confidence regions, uniformly-most-powerful tests.
  - Frequentist statistics in GW astronomy: false alarm rates, Fisher Matrix, PSD estimation.
  - Stochastic processes, optimal filtering, signal-to-noise ratio, sensitivity curves.
  - Practical: simulating random variables in python.

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# Outline of course

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- ❖ **Part 2 (lectures 7 to 12): Bayesian statistics**
  - Bayes' theorem, conjugate priors, Jeffrey's prior.
  - Bayesian hypothesis testing, hierarchical models, posterior predictive checks.
  - Sampling methods for Bayesian inference.
  - Bayesian statistics in GW astronomy: parameter estimation, population inference, model selection.
  - Practicals (2): sampling posterior distributions using pyMC3.

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# Outline of course

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- ❖ **Part 3 (lectures 13 to 16): Introduction to machine learning**
  - Introduction to machine learning.
  - Neural networks and deep learning.
  - Machine learning for GW astronomy.
  - Practical: GW search and PE using machine learning.

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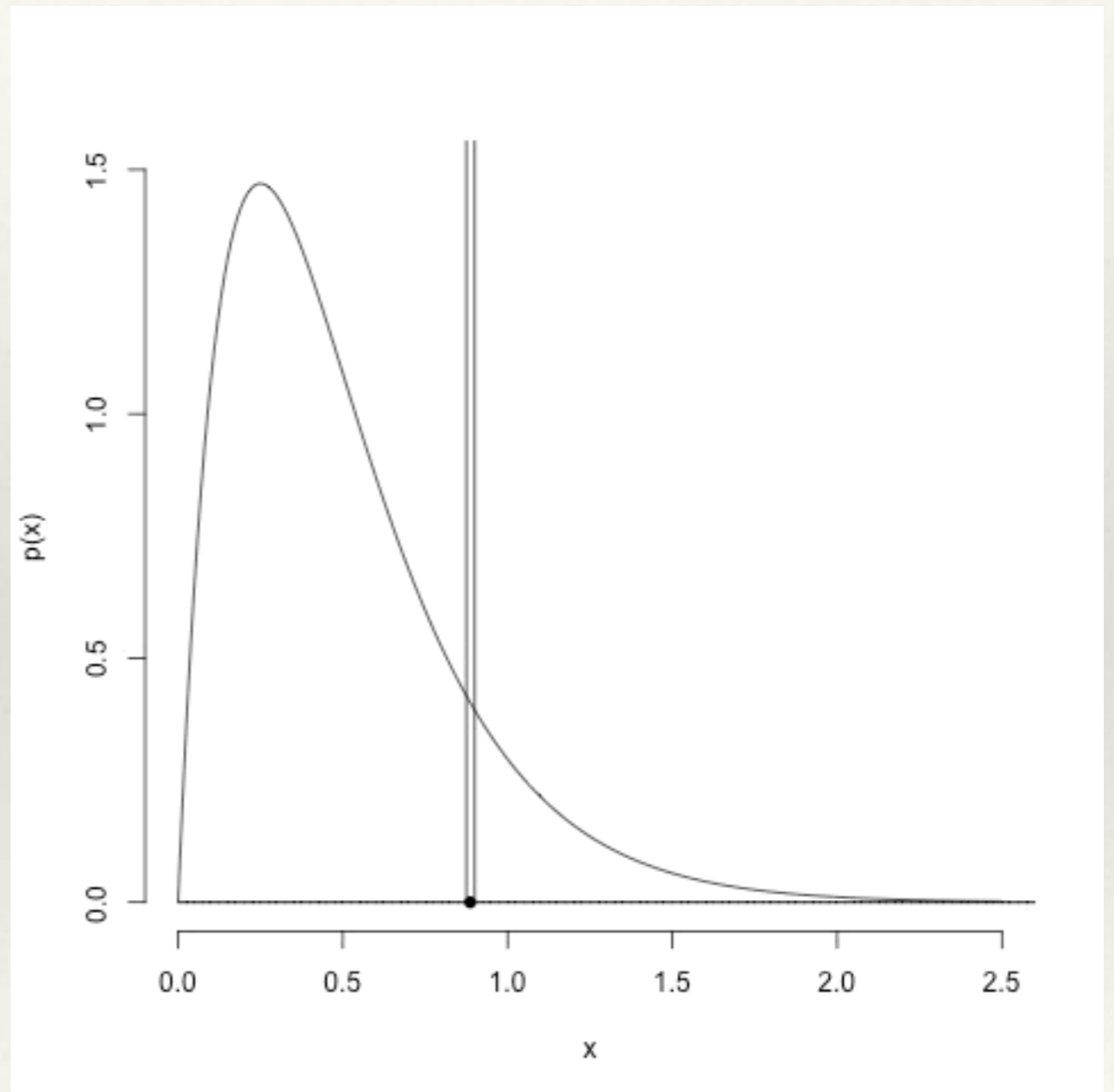
# Outline of course

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- ❖ Lecture notes will be made available on the course website
  - <https://imprs-gw-lectures.aei.mpg.de/2023-making-sense-of-data/>
- ❖ These notes will include more material than will be covered in lectures, and these extra topics will be denoted by an asterix.
- ❖ Two problem sets will be provided (one for each of the first two parts of the course). Solutions will be made available later.
- ❖ Problem sheets will have two parts, with the second part containing optional questions that are either more difficult or similar to questions in the first part.

# Random variables

- ❖ *Random variables* are quantities that are not fixed, but can take new values each time they are observed (a *realisation*).
- ❖ Over many realisations the distribution of the random variable is described by a *probability distribution*.
- ❖ Random variables can be *discrete* (taken values in a countable set) or *continuous* (taking real values in some interval).





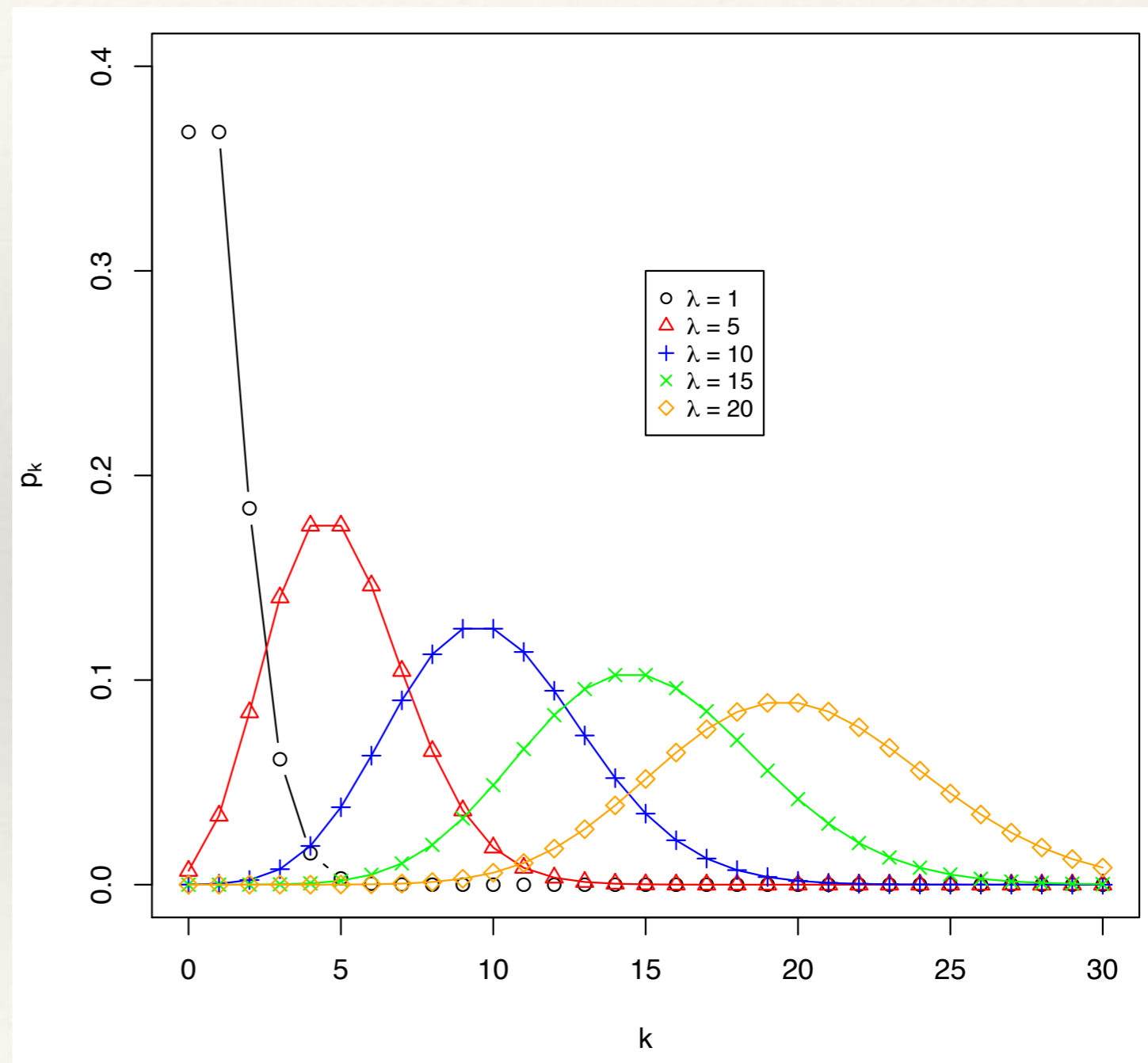


# Discrete RVs: Poisson distribution

- ❖ **Poisson distribution** is defined for non-negative  $k$  by

$$P(X = k) = p_k = \begin{cases} \lambda^k e^{-\lambda} / k! \\ 0 \end{cases}$$

- ❖ Arises as the distribution of the number of counts of a process occurring in a certain period of time.



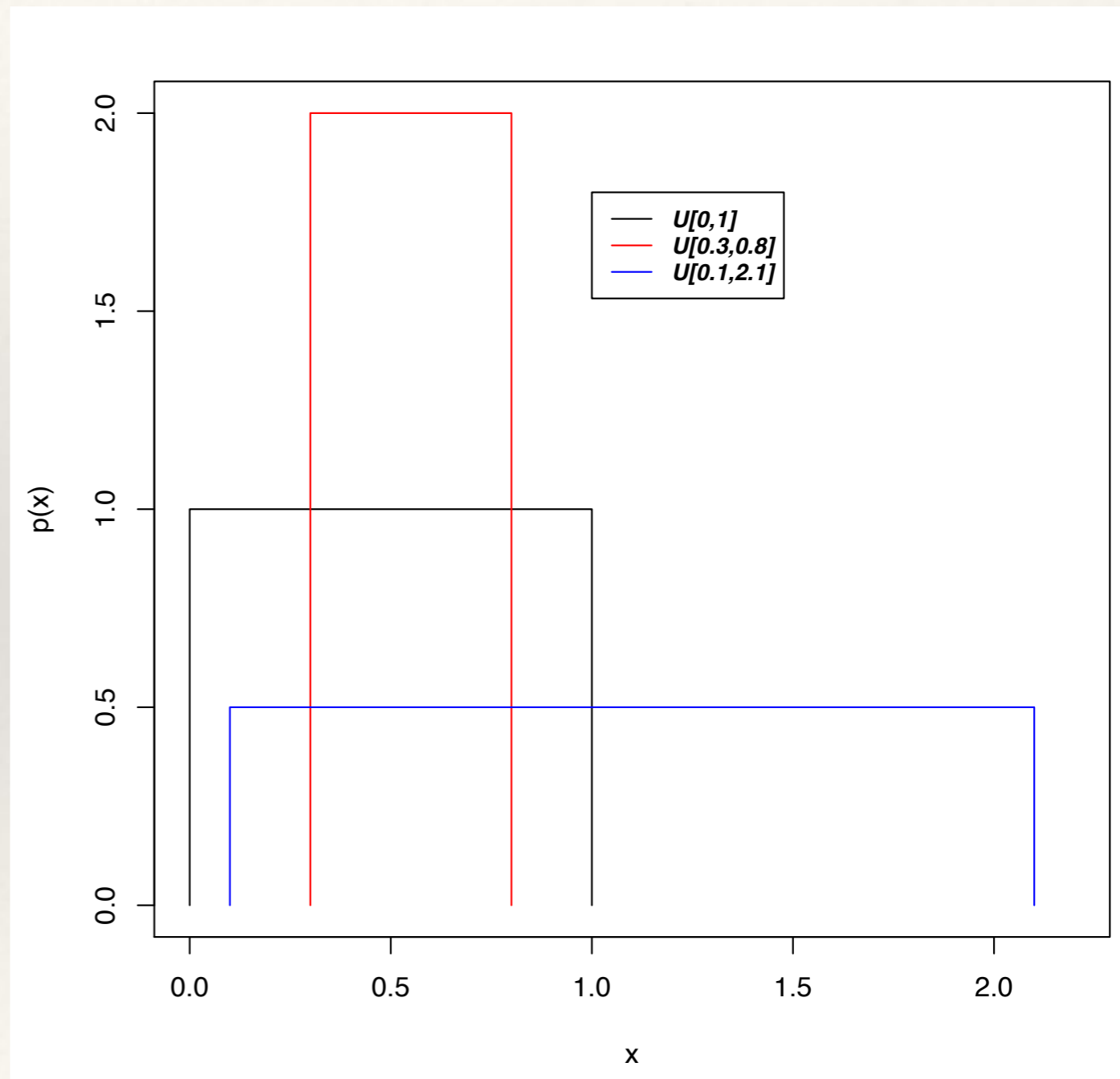
# Continuous random variables

❖ Continuous random variables are characterised by a *probability density function*, satisfying

$$0 \leq p(x) \quad \int_{x \in \mathcal{X}} p(x) dx = 1$$

❖ For example, **Uniform distribution**

$$p(x) = \begin{cases} \frac{1}{b-a} & \text{if } x \in [a, b] \\ 0 & \text{otherwise} \end{cases}$$

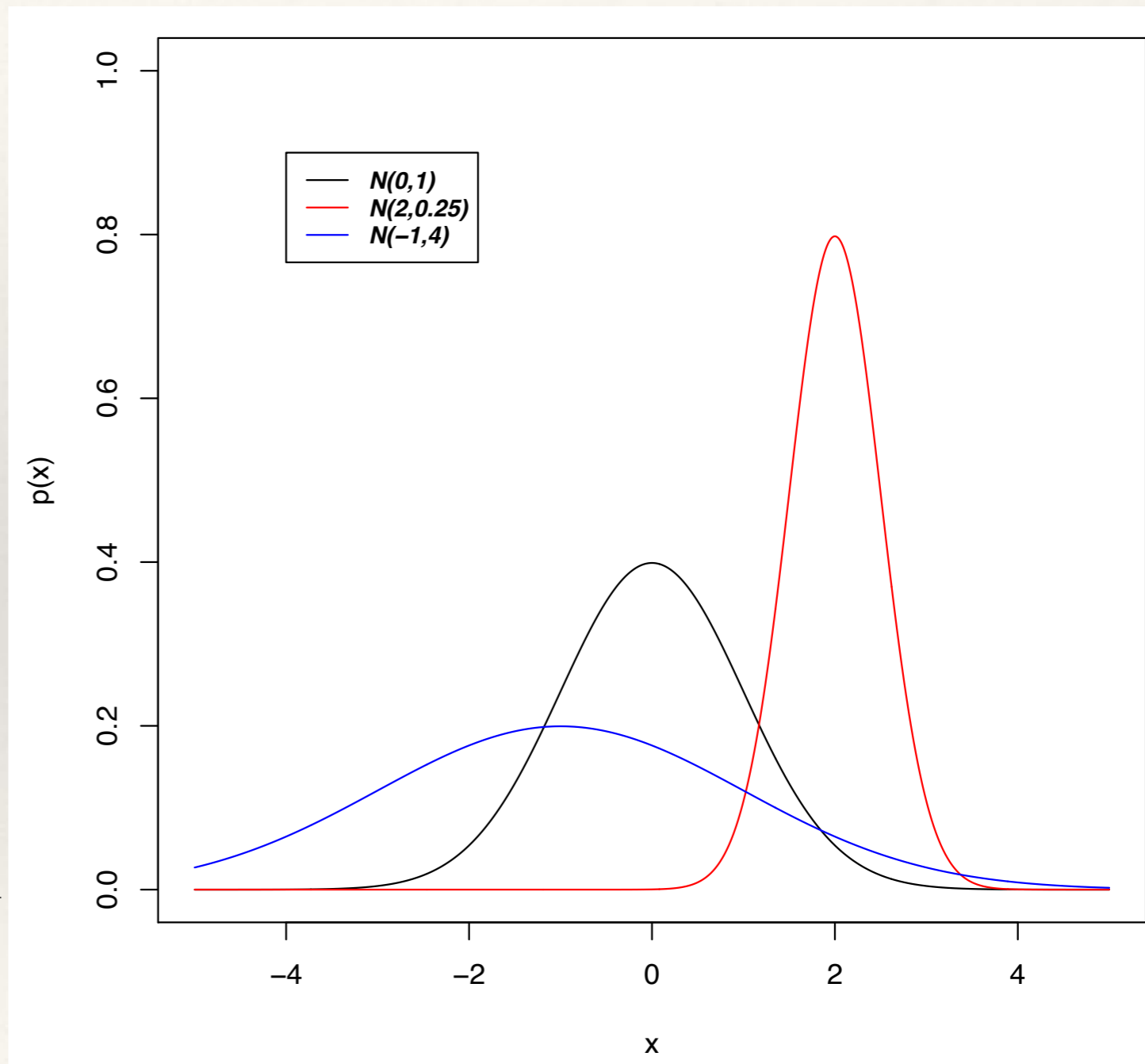


# Continuous RVs: Normal distribution

- ❖ **Normal distribution** is characterised by mean  $\mu$  and variance  $\sigma^2$

$$p(x) = \frac{1}{\sqrt{2\pi}\sigma} \exp\left(-\frac{(x - \mu)^2}{2\sigma^2}\right)$$

- ❖ Arises as a limiting distribution and as the distribution of noise in gravitational wave detectors. Commonly used as the default distribution in parametric statistics and as a prior in Bayesian analysis.
- ❖ Normal distribution with zero mean and unit variance is the *standard Normal distribution*.

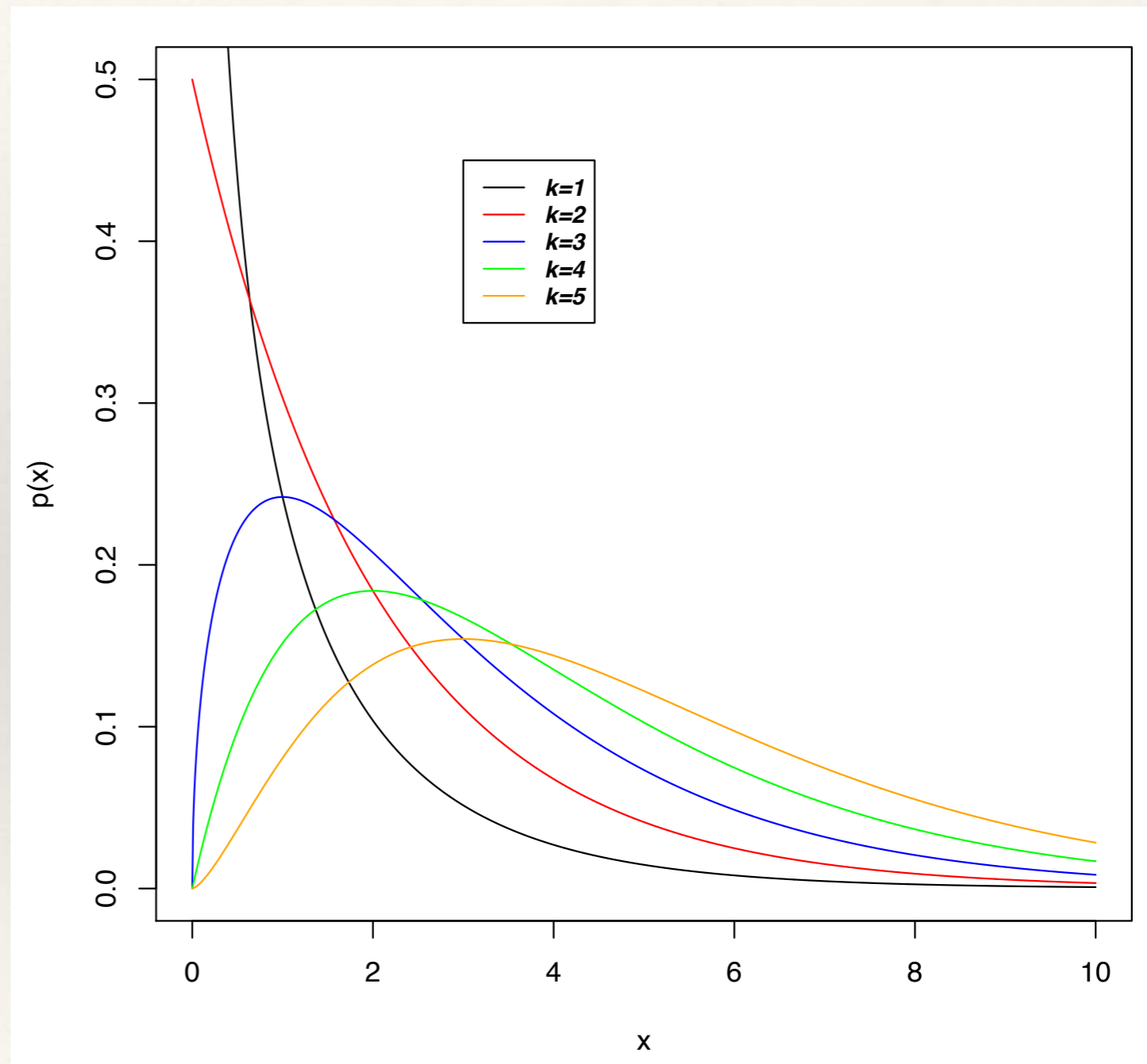


# Continuous RVs: chi-squared distribution

- ❖ **Chi-squared distribution** depends on a *degrees of freedom* parameter  $k > 0$

$$p(x) = \frac{1}{2^{k/2}\Gamma(k/2)} x^{\frac{k}{2}-1} e^{-\frac{x}{2}}$$

- ❖ It is the distribution of the sum of squares of  $k$  standard normal random variables.
- ❖ There is also a *non-central chi-square distribution* which has also a *non-centrality parameter*.

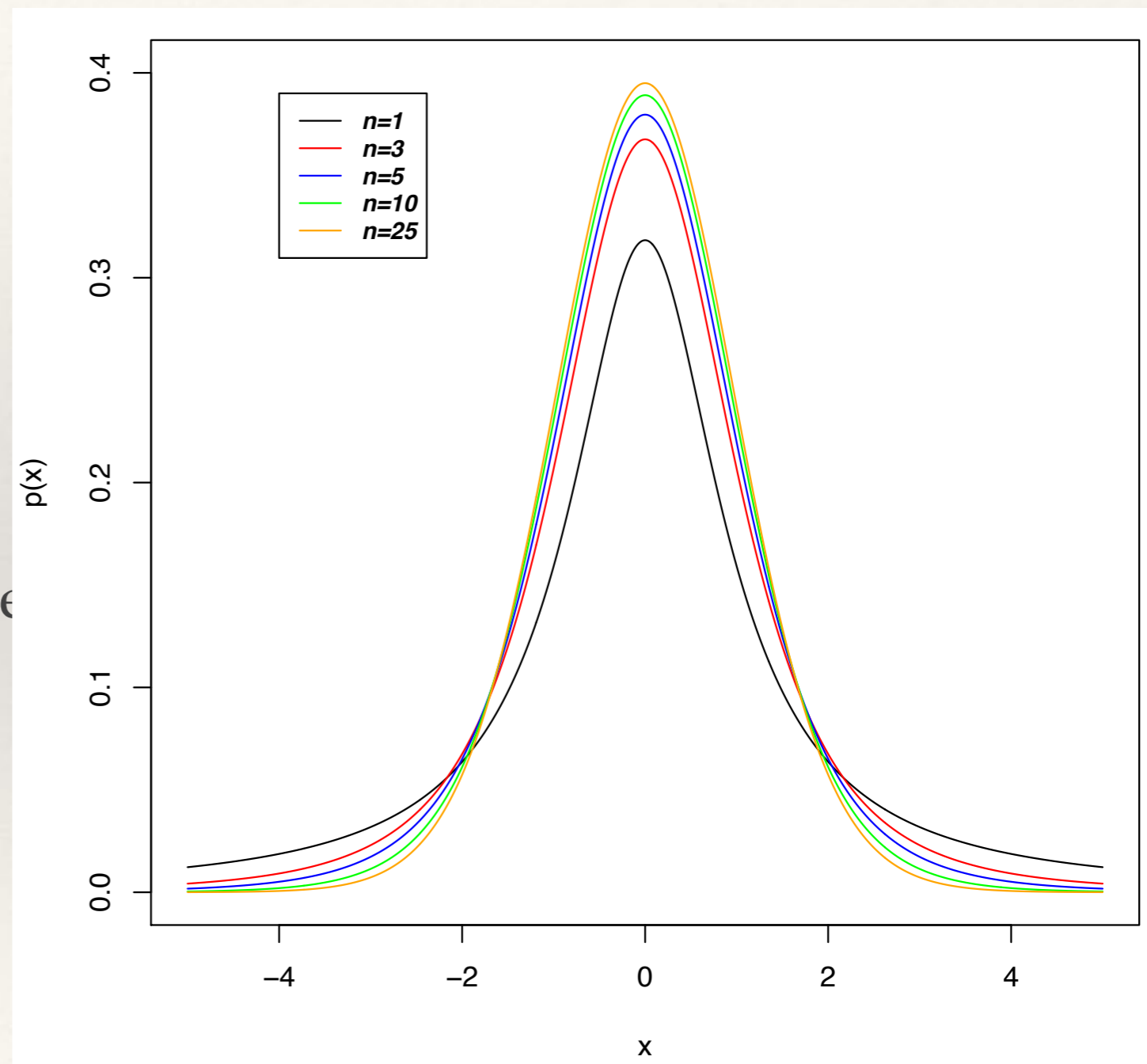


# Continuous RVs: Student's t-distribution

- ❖ Student's t-distribution *also* depends on a degrees of freedom parameter  $n$

$$p(x) = \frac{\Gamma\left(\frac{n+1}{2}\right)}{\sqrt{n\pi}\Gamma\left(\frac{n}{2}\right)} \left(1 + \frac{x^2}{n}\right)^{-\frac{n+1}{2}}$$

- ❖ It arises in hypothesis testing as the ratio of a standard Normal distribution to a chi-squared distribution. It is used as a *heavy-tailed* distribution in inference.

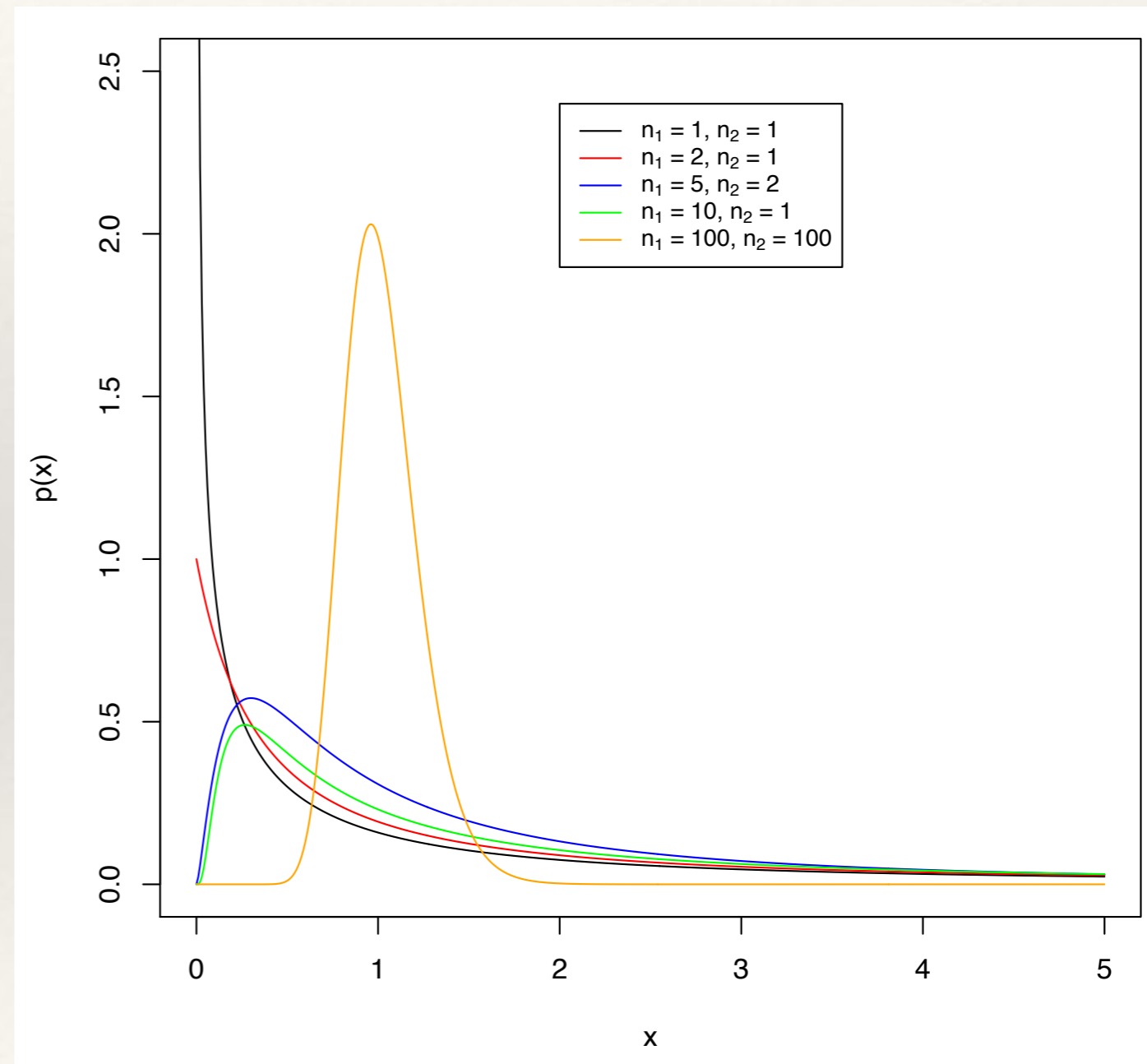


# Continuous RVs: F-distribution

- ❖ The **F-distribution** depends on two degrees of freedom parameters,  $n_1$  and  $n_2$

$$p(x) = \frac{1}{B\left(\frac{n_1}{2}, \frac{n_2}{2}\right)} \left(\frac{n_1}{n_2}\right)^{\frac{n_1}{2}} x^{\frac{n_1}{2}-1} \left(1 + \frac{n_1}{n_2}x\right)^{-\frac{n_1+n_2}{2}}$$

- ❖ This arises as the ratio of two chi-square distributions and is the basis for *analysis of variance*.

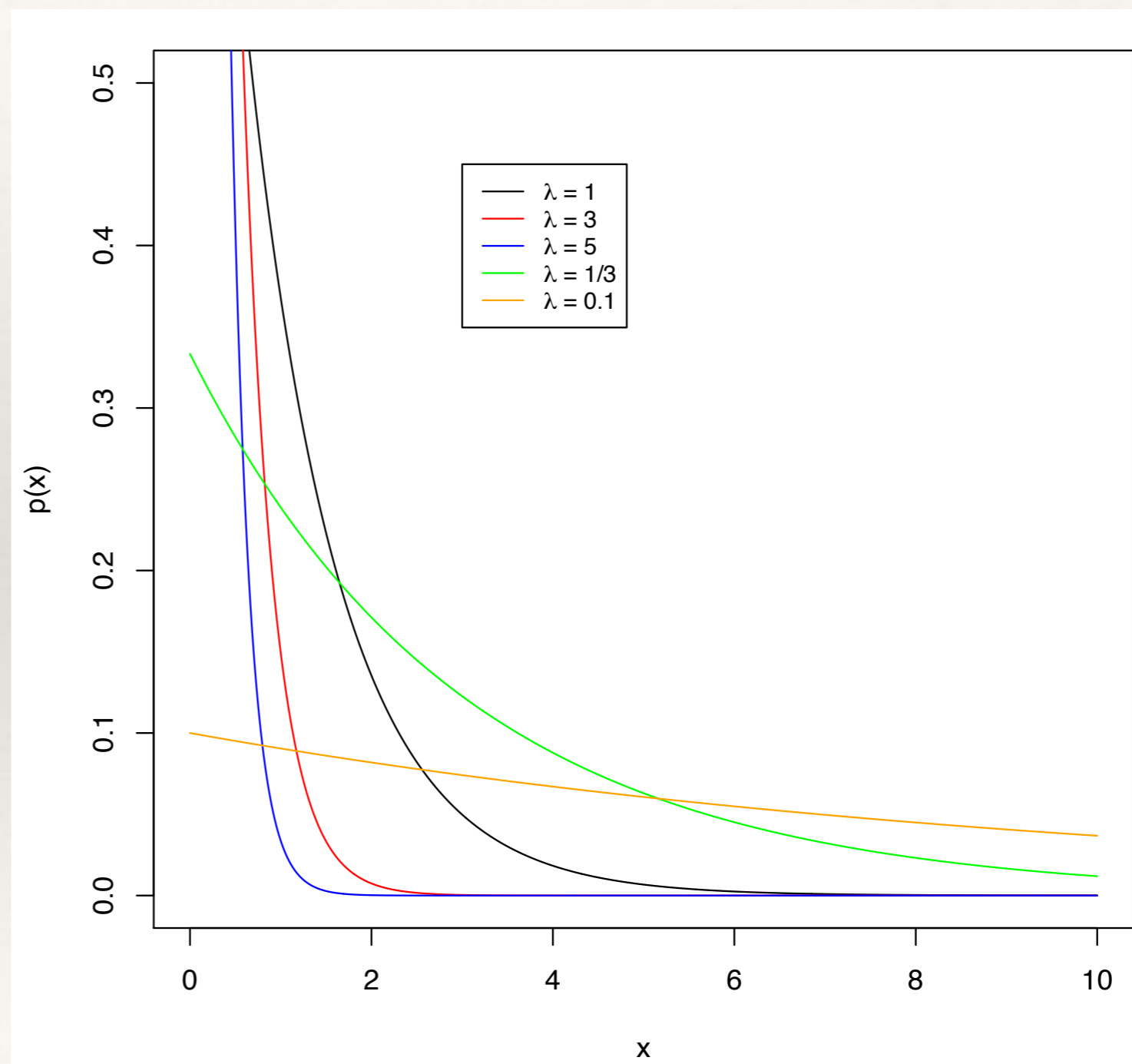


# Continuous RVs: Exponential distribution

- ❖ The **Exponential distribution** depends on a *rate* parameter  $\lambda > 0$

$$p(x) = \begin{cases} \lambda e^{-\lambda x} & \text{if } x > 0 \\ 0 & \text{otherwise} \end{cases}$$

- ❖ This arises as the distribution of the separation of events in a Poisson process.



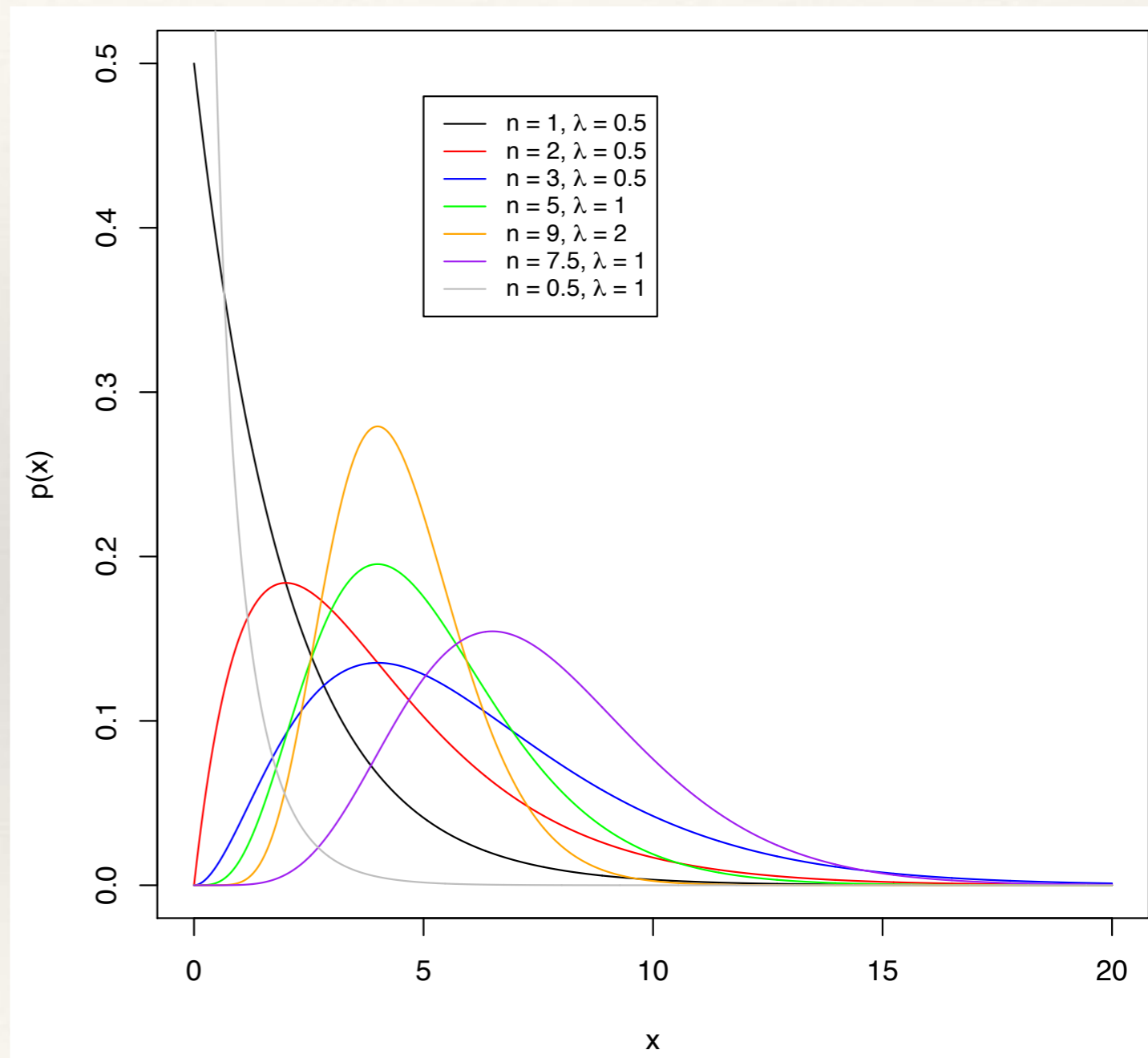


# Continuous RVs: Gamma distribution

- ❖ The **Gamma distribution** depends on a *shape parameter*  $n > 0$  and a *scale parameter*  $\lambda > 0$

$$p(x) = \begin{cases} \frac{1}{\Gamma(n)} \lambda^n x^{n-1} e^{-\lambda x} & \text{if } x > 0 \\ 0 & \text{otherwise} \end{cases}$$

- ❖ The Gamma distribution is commonly used in Bayesian inference as a prior with support on the positive real line, and is conjugate to the Poisson distribution.



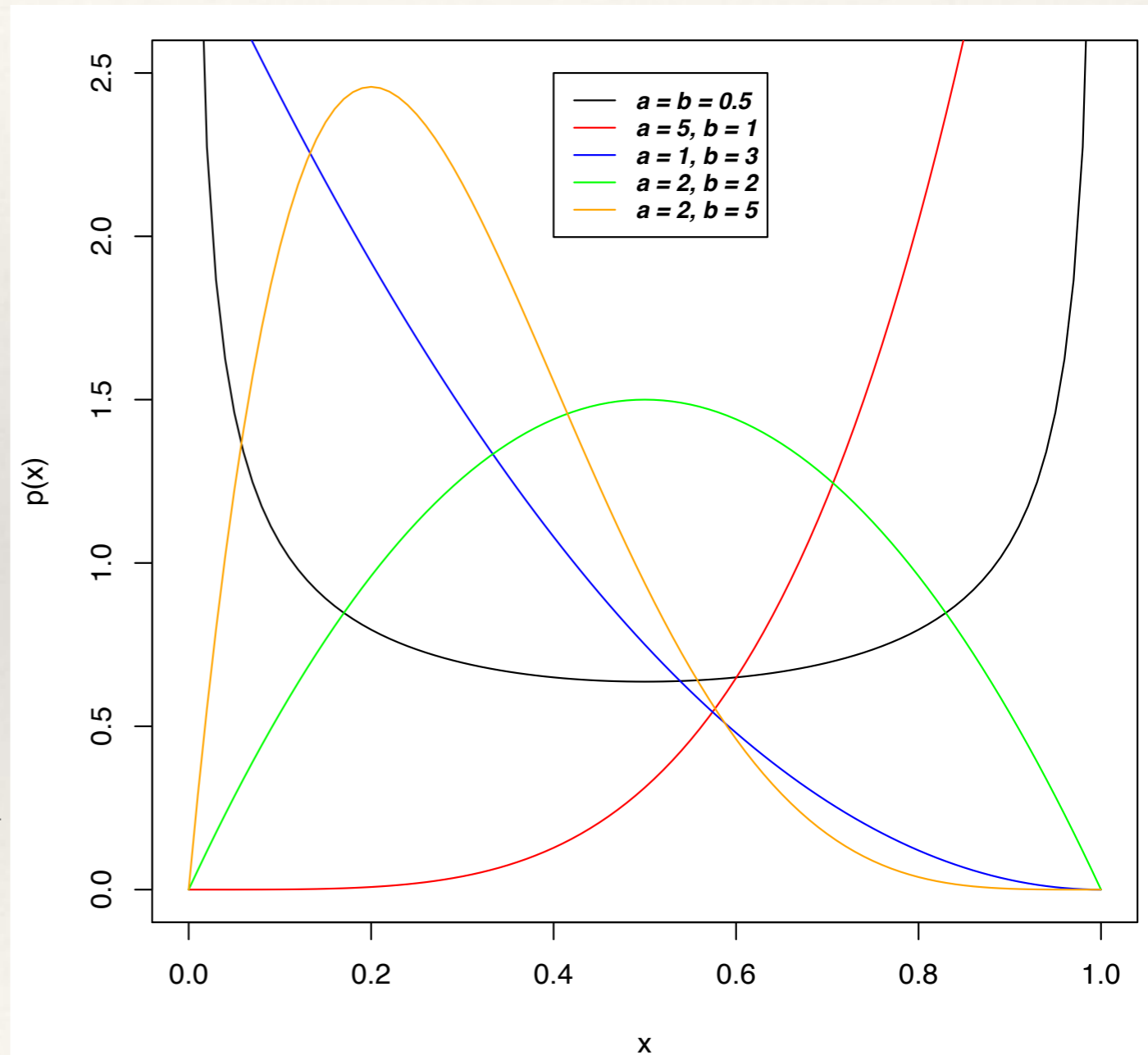
# Continuous RVs: Beta distribution

- ❖ The **Beta distribution** depends on two *shape parameters*  $a, b > 0$

$$p(x) = \begin{cases} \frac{1}{B(a,b)} x^{a-1} (1-x)^{b-1} & \text{if } 0 < x < 1 \\ 0 & \text{otherwise} \end{cases}$$

$$B(a, b) = \int_0^1 x^{a-1} (1-x)^{b-1} dx$$

- ❖ The Beta distribution is conjugate to the Binomial distribution and is used as a prior for parameters with support in  $[0,1]$ .



# Continuous RVs: Dirichlet distribution

- ❖ The **Dirichlet distribution** is a *multivariate distribution*, generating  $K$  samples  $\{x_i\}$  constrained such that  $0 < x_i < 1$  and

$$\sum_{i=1}^K x_i = 1$$

- ❖ The distribution depends on a vector of *concentration parameters*

$$\vec{\alpha} = (\alpha_1, \dots, \alpha_K)$$

- ❖ and has pdf

$$p(x) = \frac{1}{B(\vec{\alpha})} \prod_{i=1}^K x_i^{\alpha_i - 1}, \quad \text{where } B(\vec{\alpha}) = \frac{\prod_{i=1}^K \Gamma(\alpha_i)}{\Gamma\left(\sum_{j=1}^K \alpha_j\right)}.$$

- ❖ The **Dirichlet process** is used as a prior on probability distributions in Bayesian nonparametric inference.

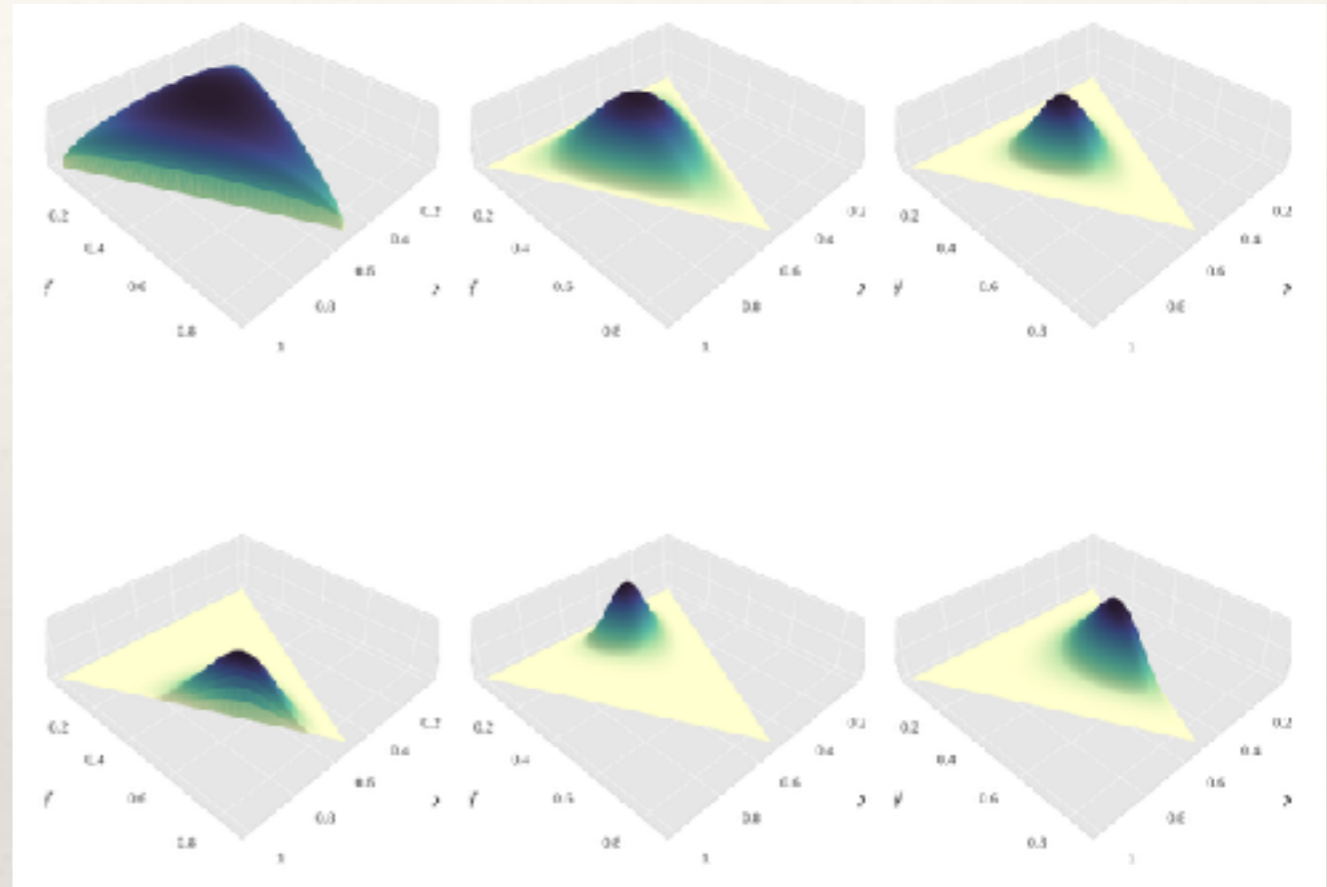


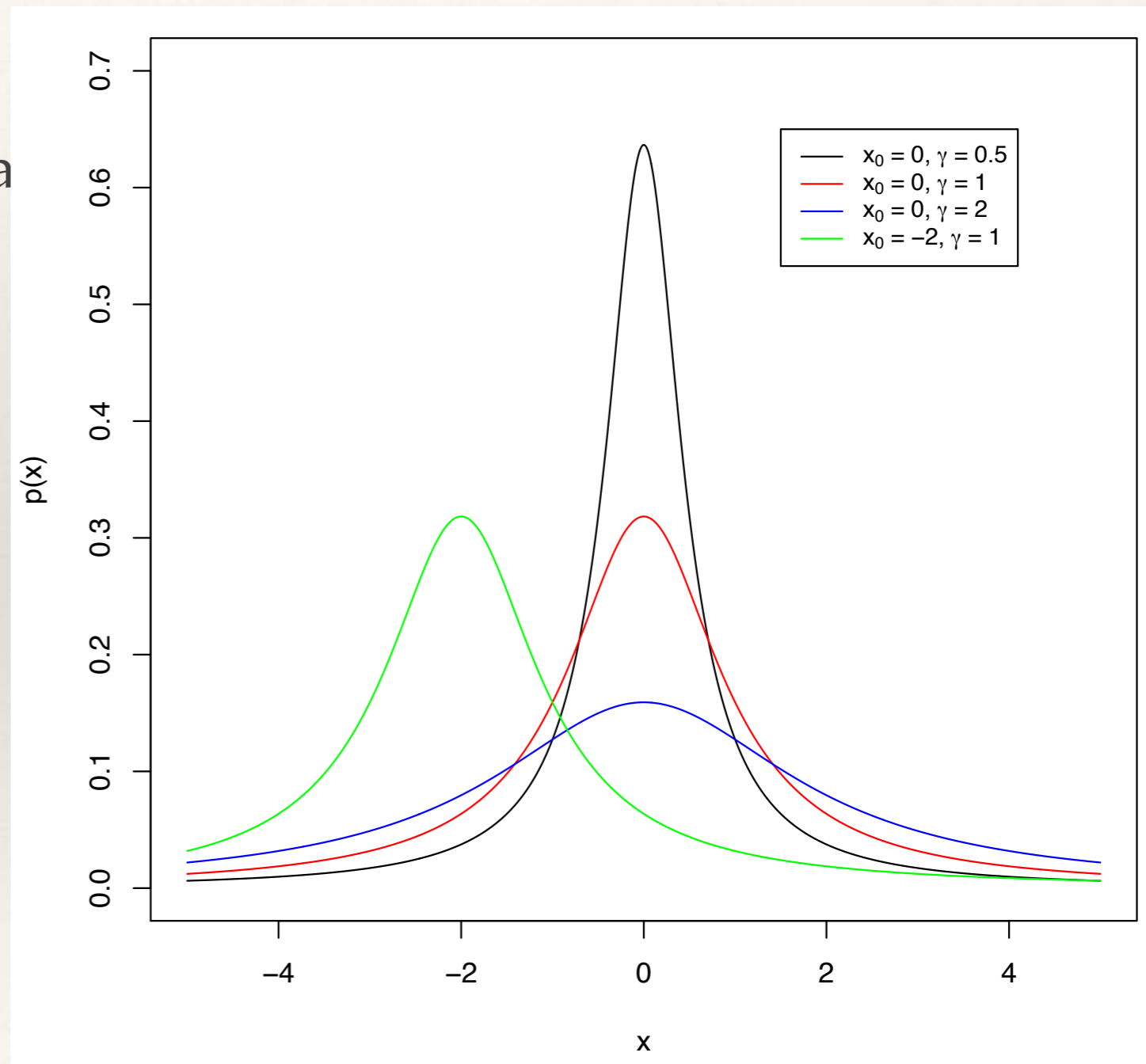
Figure from *Wikipedia*

# Continuous RVs: Cauchy distribution

- ❖ The **Cauchy distribution** (or Lorentz distribution) depends on a *location parameter*,  $x_0$ , and a *scale parameter*,  $\gamma > 0$

$$p(x) = \frac{1}{\pi\gamma \left[ 1 + \left( \frac{x-x_0}{\gamma} \right)^2 \right]}$$

- ❖ This distribution arises in optics and is used to model distributions with sharp features, e.g., spectral lines in LIGO.



# Summarising random variables: average

❖ The pdf (or pmf) completely characterises a probability distribution, but it is often more convenient to work with summary quantities.

❖ These are based on *expectation values*

$$\mathbb{E}(T(X)) = \int_{-\infty}^{\infty} p(x)t(x)dx$$

❖ There are various quantities that summarise the *average* value of a random variable

- **Mean**

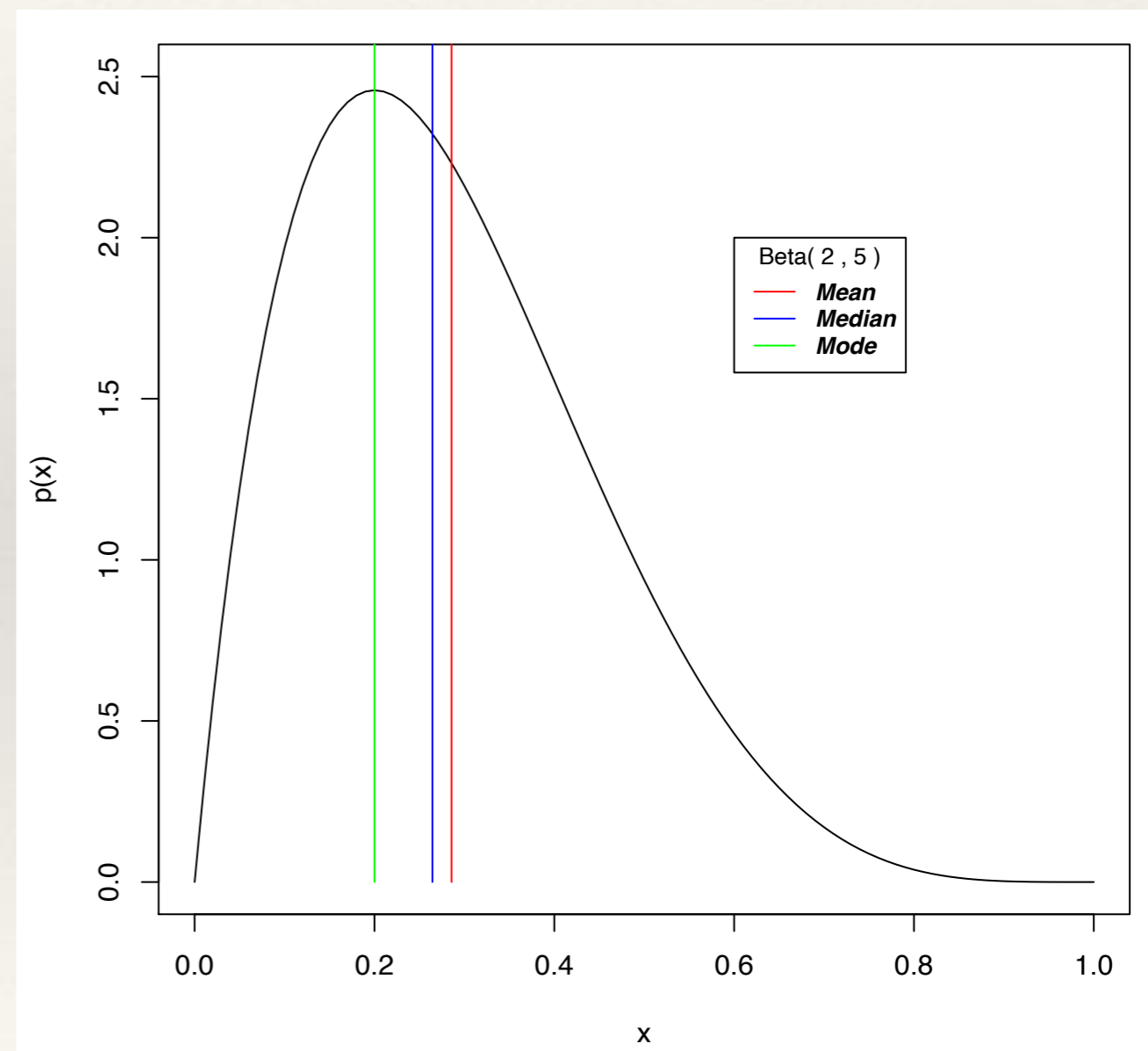
$$\mu = \mathbb{E}(X)$$

- **Median**  $m$  satisfies

$$\int_{-\infty}^m p(x)dx = \int_m^{\infty} p(x)dx = \frac{1}{2}$$

- **Mode**

$$M = \operatorname{argmax}_{x \in \mathcal{X}} p(x)$$

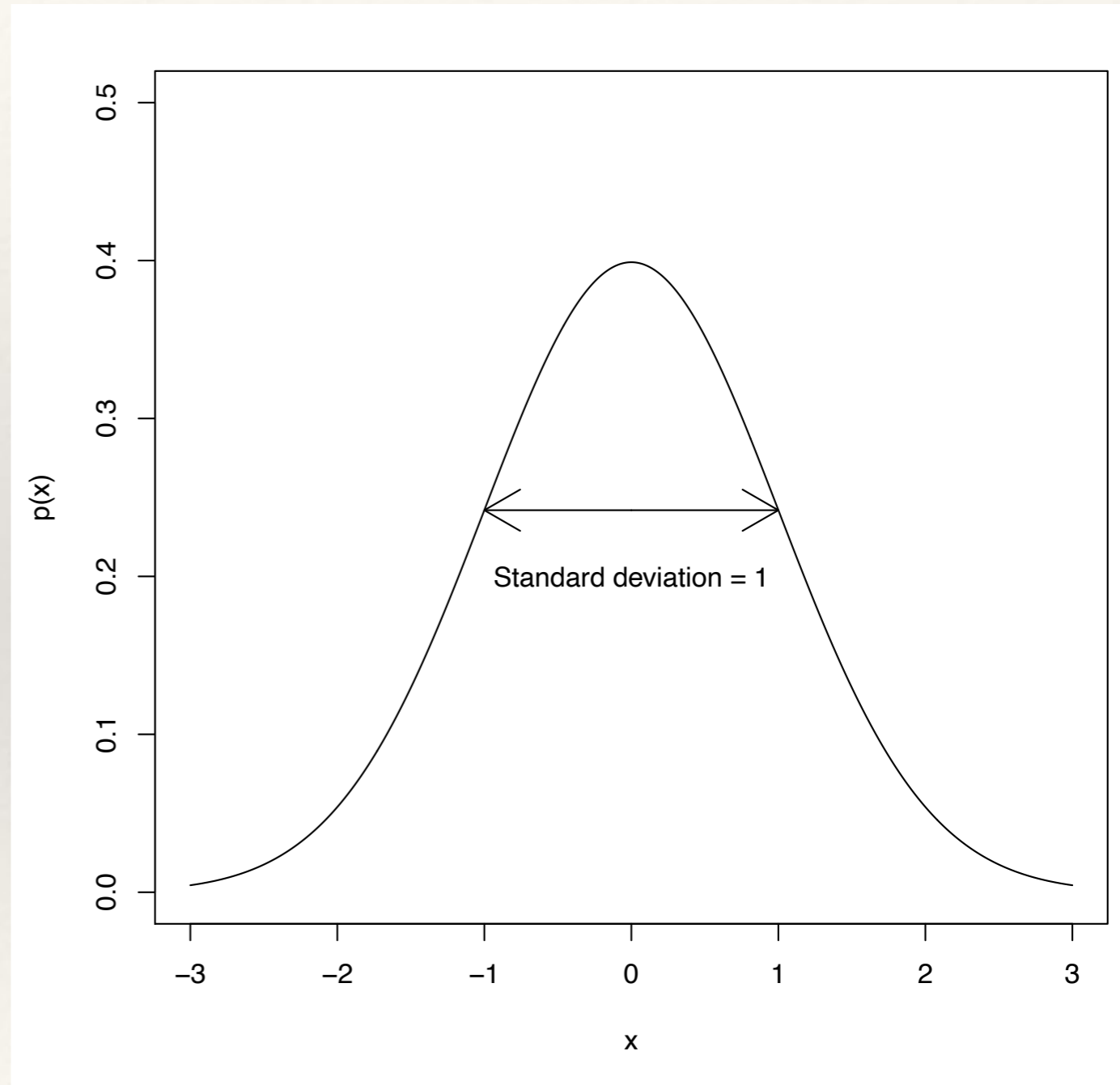


# Summarising random variables: spread

❖ Other quantities summarise the spread of a RV

- **Variance/Standard deviation**

$$\text{Var}(X) = \mathbb{E} [(X - \mathbb{E}(X))^2]$$



# Summarising random variables: spread

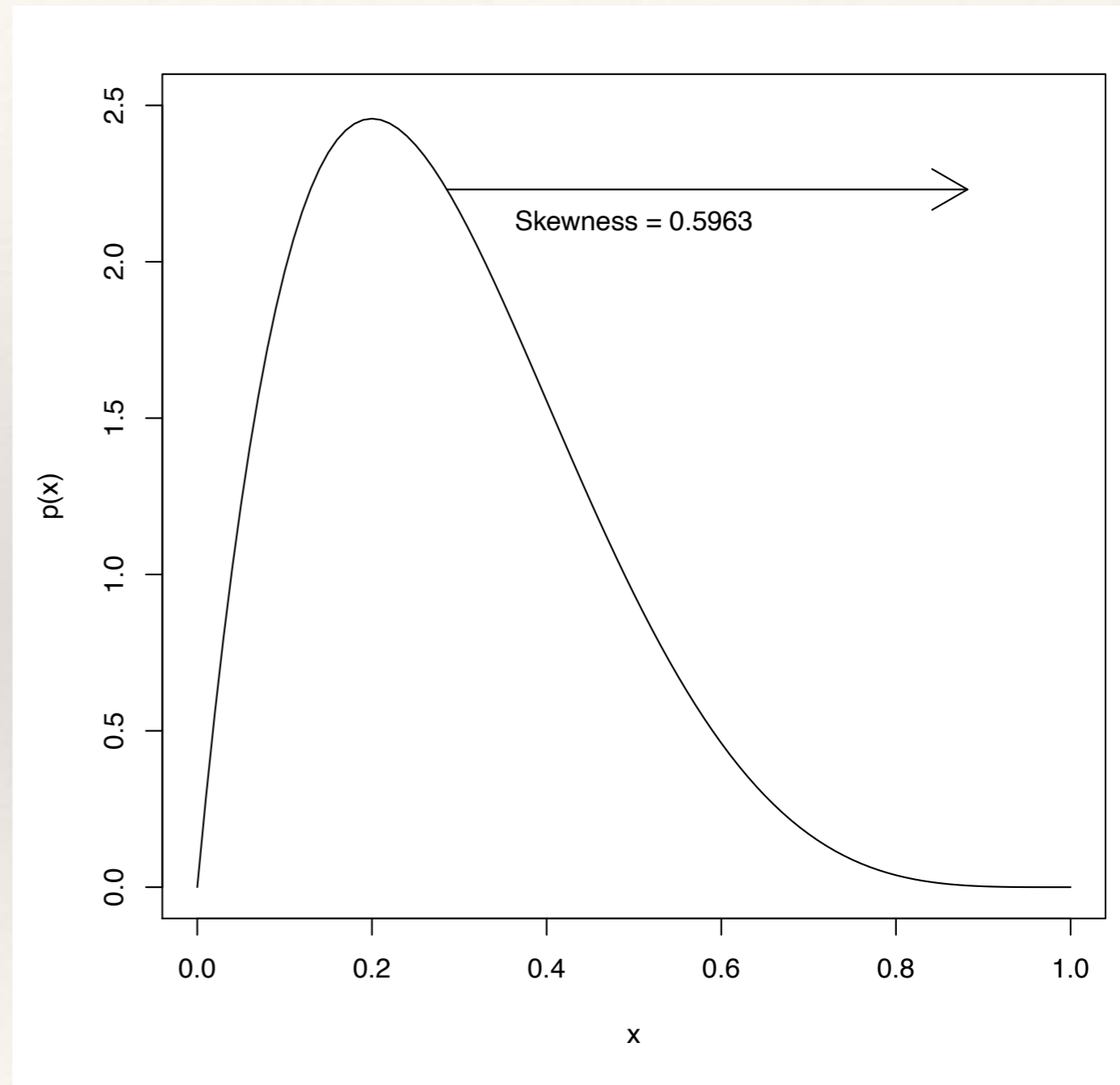
❖ Other quantities summarise the spread of a RV

- **Variance/Standard deviation**

$$\text{Var}(X) = \mathbb{E} \left[ (X - \mathbb{E}(X))^2 \right]$$

- **Skewness**

$$\gamma_1 = \mathbb{E} \left[ \left( \frac{x - \mu}{\sigma} \right)^3 \right]$$



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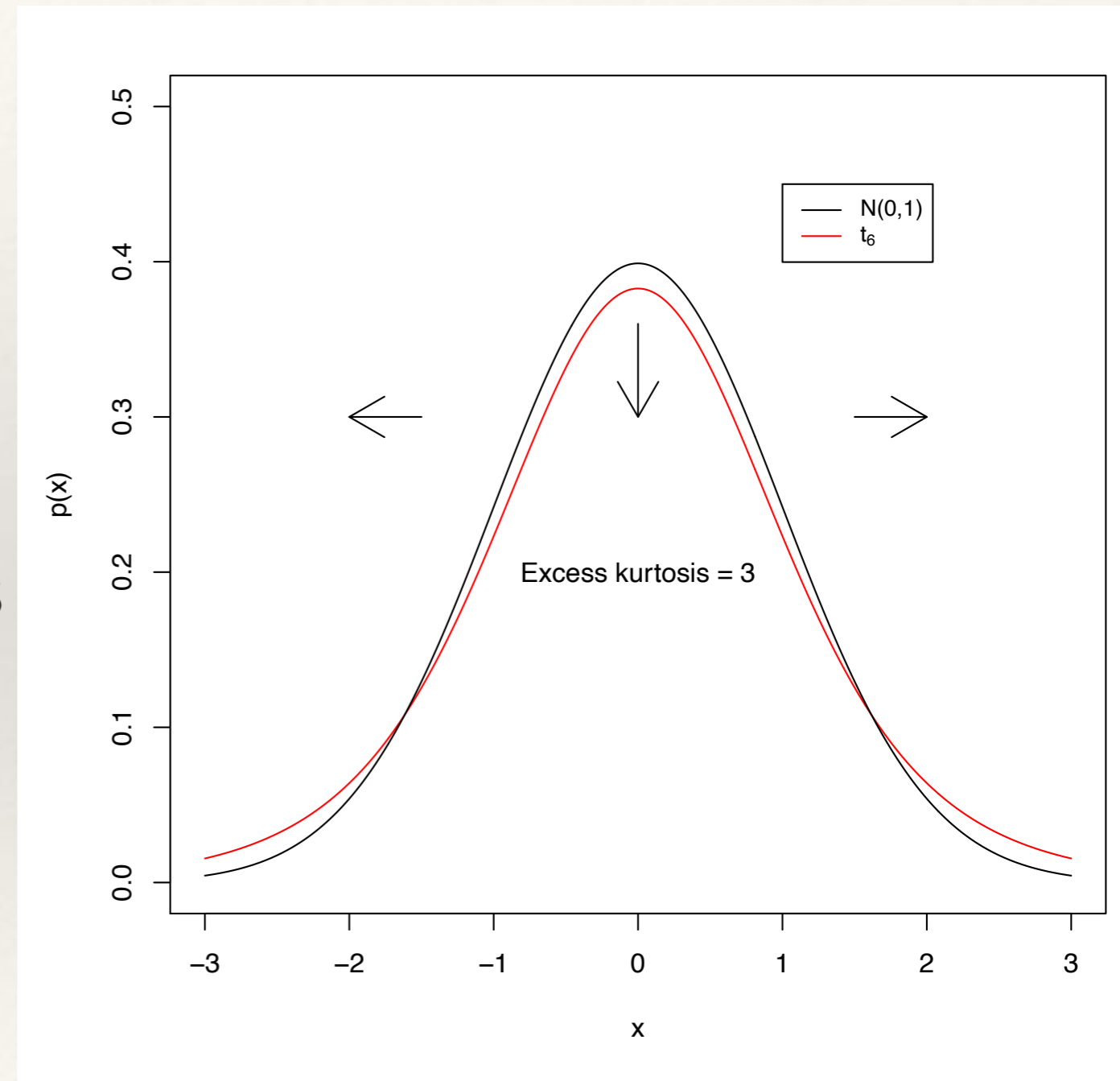
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- **Excess Kurtosis**

$$\text{Kurt}(X) = \mathbb{E} \left[ \left( \frac{x - \mu}{\sigma} \right)^4 \right] - 3$$





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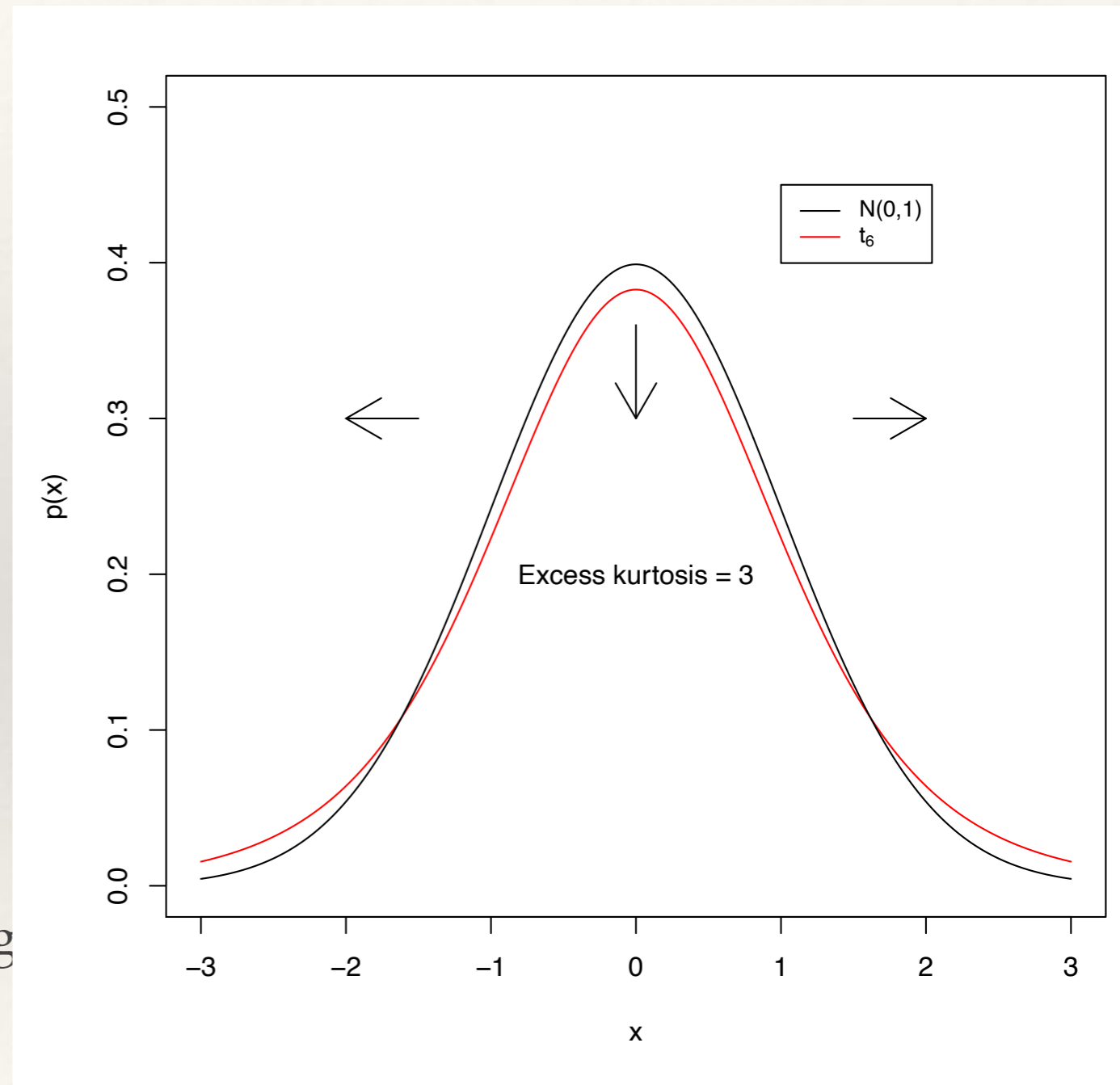
$$\text{Kurt}(X) = \mathbb{E} \left[ \left( \frac{x - \mu}{\sigma} \right)^4 \right] - 3$$

- **Higher moments**

$$\mathbb{E} \left[ (X - c)^n \right]$$

❖ Moments can be efficiently computed using the *moment generating function*

$$M_X(t) = \mathbb{E} \left[ e^{tX} \right] \quad t \in \mathbb{R}$$



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# Independence

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- ❖ A set of random variables  $\{X_1, X_2, \dots, X_N\}$  is *independent* if, for all choices of  $\{x_1, x_2, \dots, x_N\}$

$$P(X_1 \leq x_1, X_2 \leq x_2, \dots, X_N \leq x_N) = P(X_1 \leq x_1)P(X_2 \leq x_2) \dots P(X_N \leq x_N)$$

- ❖ In terms of the density function this is equivalent to

$$p(x_1, \dots, x_N) = p_{X_1}(x_1)p_{X_2}(x_2) \dots p_{X_N}(x_N)$$

- ❖ Two independent random variables have zero covariance

$$\text{cov}(X, Y) = \mathbb{E} [(X - \mathbb{E}(X)) (Y - \mathbb{E}(Y))] = 0$$

- ❖ but the converse is not necessarily true.
- ❖ Random variables are *independent identically distributed* (IID) if they are independent and are all drawn from the same probability distribution.

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# Linear combinations of RVs

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- ❖ Suppose  $X_1, \dots, X_N$  are random variables and consider a new RV

$$Y = \sum_{i=1}^N a_i X_i$$

- ❖  $Y$  has the properties

$$\mathbb{E}(Y) = \sum_{i=1}^N a_i \mathbb{E}(X_i), \quad \text{Var}(Y) = \sum_{i=1}^N a_i^2 \text{Var}(X_i) + \sum_{i \neq j} a_i a_j \text{cov}(X_i, X_j)$$

- ❖ The first equation holds for any random variables. If the RVs are *independent* then the relationships simplify

$$\text{Var}(Y) = \sum_{i=1}^N a_i^2 \text{Var}(X_i) \quad M_Y(t) = \prod_{i=1}^N M_{X_i}(a_i t)$$

- ❖ If  $\{X_i\}$  are IID then the *sample mean* defined by  $a_i=1/N$  for all  $i$  has the properties

$$\mathbb{E}(\hat{\mu}) = \mathbb{E}(X_1), \quad \text{Var}(\hat{\mu}) = \frac{1}{n} \text{Var}(X_1), \quad M_{\hat{\mu}}(t) = \left( M_{X_1} \left( \frac{t}{N} \right) \right)^N$$

# Laws of large numbers

- ❖ Averages of random variables have various nice asymptotic properties

$$S_n = \sum_{i=1}^n X_i \quad \mathbb{E}(X) = \mu \quad \text{Var}(X) = \sigma^2$$

- ❖ *Weak law of large numbers*: for  $\epsilon > 0$

$$P\left(\left|\frac{S_n}{n} - \mu\right| > \epsilon\right) \rightarrow 0, \text{ as } n \rightarrow \infty$$

- ❖ *Strong law of large numbers*

$$P\left(\frac{S_n}{n} \rightarrow \mu\right) = 1$$

- ❖ *Central Limit Theorem*: for  $S_n^* = \frac{S_n - n\mu}{\sigma\sqrt{n}}$

$$\lim_{n \rightarrow \infty} P(a \leq S_n^* \leq b) = \Phi(b) - \Phi(a)$$

